

ENEE630 Part-3

## ***Part 3. Spectrum Estimation***

### ***3.2 Parametric Methods for Spectral Estimation***

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Acknowledgment: ENEE630 slides were based on class notes developed by Profs. K.J. Ray Liu and Min Wu. The slides were made by Prof. Min Wu, with updates from Mr. Wei-Hong Chuang. *Contact: [minwu@eng.umd.edu](mailto:minwu@eng.umd.edu)*



# **Summary of Related Readings on Part-III**

Overview Haykins 1.16, 1.10

## 3.1 Non-parametric method

Hayes 8.1; 8.2 (8.2.3, 8.2.5); 8.3

## 3.2 Parametric method

Hayes 8.5, 4.7; 8.4

## 3.3 Frequency estimation

Hayes 8.6

## Review

- On DSP and Linear algebra: Hayes 2.2, 2.3
- On probability and parameter estimation: Hayes 3.1 – 3.2

# **Motivation**

- **Implicit assumption by classical methods**
  - Classical methods use Fourier transform on either windowed data or windowed autocorrelation function (ACF)
  - Implicitly assume the unobserved data or ACF outside the window are zero => not true in reality
  - Consequence of windowing: smeared spectral estimate (leading to low resolution)
- **If prior knowledge about the process is available**
  - We can use prior knowledge and select a good model to approximate the process
  - Usually need to estimate fewer model parameters (than non-parametric approaches) using the limited data points we have
  - The model may allow us to better describe the process outside the window (instead of assuming zeros)

# **General Procedure of Parametric Methods**

- **Select a model** (based on prior knowledge)
- **Estimate the parameters of the assumed model**
- **Obtain the spectral estimate implied by the model** (with the estimated parameters)

## **Spectral Estimation using AR, MA, ARMA Models**

- **Physical insight:** the process is generated/approximated by filtering white noise with an LTI filter of rational transfer func  $H(z)$
- **Use observed data to estimate a few lags of  $r(k)$** 
  - Larger lags of  $r(k)$  can be implicitly extrapolated by the model
- **Relation between  $r(k)$  and filter parameters  $\{a_k\}$  and  $\{b_k\}$** 
  - PARAMETER EQUATIONS from Section 2.1.2(6)
  - Solve the parameter equations to obtain filter parameters
  - Use the p.s.d. implied by the model as our spectral estimate
- **Deal with nonlinear parameter equations**
  - Try to convert/relate them to the AR models that have linear equations

# Review: Parameter Equations

Yule-Walker equations (for AR process)

$$\Gamma_x[k] = \begin{cases} -\sum_{l=1}^p a[l] \Gamma_x[-l] + \sigma^2 & \text{for } k=0 \\ -\sum_{l=1}^p a[l] \Gamma_x[k-l] & \text{for } k \geq 1 \end{cases}$$

$$\begin{bmatrix} \Gamma_x(0) & \Gamma_x(1) & \dots & \Gamma_x(-p+1) \\ \Gamma_x(1) & \Gamma_x(0) & \dots & \Gamma_x(-p+2) \\ \vdots & & \ddots & \vdots \\ \Gamma_x(p-1) & \dots & & \Gamma_x(0) \end{bmatrix} \begin{bmatrix} a[1] \\ a[2] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} \Gamma_x(1) \\ \Gamma_x(2) \\ \vdots \\ \Gamma_x(p) \end{bmatrix}$$

ARMA model

$$\Gamma_x[k] = \begin{cases} -\sum_{l=1}^p a[l] \Gamma_x[k-l] + \sigma^2 \sum_{l=0}^{q-k} h^*[l] b[l+k] & k=0, 1, \dots, q \\ -\sum_{l=1}^p a[l] \Gamma_x[k-l] & k \geq q+1 \end{cases}$$

MA model

$$\Gamma_x[k] = \begin{cases} \sigma^2 \sum_{l=0}^{q-k} b^*[l] b[l+k] & \text{for } k=0, 1, \dots, q \\ 0 & \text{for } k \geq q+1 \end{cases}$$

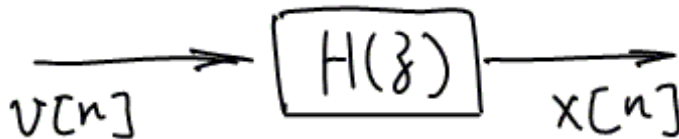
## 3.2.1 AR Spectral Estimation

### (1) Review of AR process

- The time series  $\{x[n], x[n-1], \dots, x[n-m]\}$  is a realization of an AR process of order  $M$  if it satisfies the difference equation

where  $\{v[n]\}$  is a white noise process with variance  $\sigma^2$ .

- Generating an AR process with parameters  $\{a_i\}$ :



## 3.2.1 AR Spectral Estimation

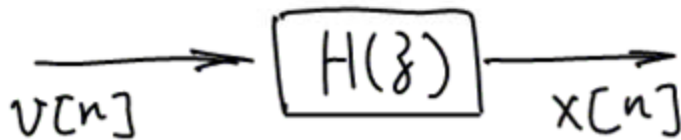
### (1) Review of AR process

- The time series  $\{x[n], x[n-1], \dots, x[n-M]\}$  is a realization of an AR process of order  $M$  if it satisfies the difference equation

$$x[n] + a_1 x[n-1] + \dots + a_M x[n-M] = v[n]$$

where  $\{v[n]\}$  is a white noise process with variance  $\sigma^2$ .

- Generating an AR process with parameters  $\{a_i\}$ :



$$H(z) = \frac{1}{1 + \sum_{i=1}^M a_i z^{-i}}$$
$$\stackrel{\text{def}}{=} \frac{1}{A(z)}$$



## P.S.D. of An AR Process

Recall: the p.s.d. of an AR process  $\{x[n]\}$  is given by

$$\hat{P}_{\text{AR}}(z) =$$

$$\Downarrow z = e^{j\omega} = e^{j2\pi f}$$

$$\hat{P}_{\text{AR}}(f) =$$

## P.S.D. of An AR Process

Recall: the p.s.d. of an AR process  $\{x[n]\}$  is given by

$$\hat{P}_{\text{AR}}(z) = \frac{\sigma^2}{A(z)A^*(1/z^*)}$$
$$\Downarrow z = e^{j\omega} = e^{j2\pi f}$$
$$\hat{P}_{\text{AR}}(f) = \frac{\sigma^2}{\left| 1 + \sum_{k=1}^M a_k e^{-j2\pi f k} \right|^2}$$

## Procedure of AR Spectral Estimation

- Observe the available data points  $x[0], \dots, x[N-1]$ , and Determine the AR process order  $p$
- Estimate the autocorrelation functions (ACF)  $k=0, \dots, p$

- Solve  $\{a_i\}$  from the Yule-Walker equations or the normal equation of forward linear prediction
  - Recall for an AR process, the normal equation of FLP is equivalent to the Yule-Walker equation
- Obtain power spectrum  $P_{AR}(f)$ :

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- Estimate the autocorrelation functions (ACF)  $k=0, \dots, p$

Biased (low variance)

$$\hat{r}(k) = \frac{1}{N} \sum_{n=0}^{N-1-k} x[n+k]x^*[n]$$

Unbiased (may not non-neg.definite)

$$\hat{r}(k) = \frac{1}{N-k} \sum_{n=0}^{N-1-k} x[n+k]x^*[n]$$

- Solve  $\{a_i\}$  from the Yule-Walker equations or the normal equation of forward linear prediction
  - Recall for an AR process, the normal equation of FLP is equivalent to the Yule-Walker equation

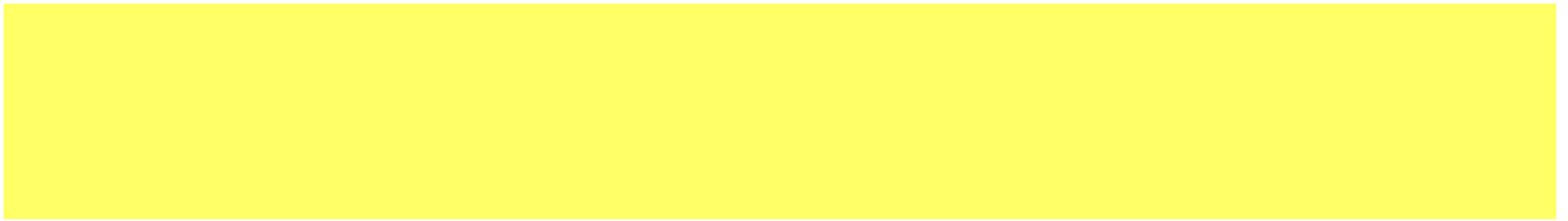
- Obtain power spectrum  $P_{AR}(f)$ : 
$$\hat{P}_{AR}(f) = \frac{\sigma^2}{\left|1 + \sum_{k=1}^M a_k e^{-j2\pi f k}\right|^2}$$

## **3.2.2 Maximum Entropy Spectral Estimation (MESE)**

- **View point: Extrapolations of ACF**
  - $\{r[0], r[1], \dots, r[p]\}$  is known; there are generally **an infinite number of possible extrapolations** for  $r(k)$  at larger lags
  - As long as  $\{r[p+1], r[p+2], \dots\}$  guarantee that the correlation matrix is non-negative definite, they all form valid ACFs for w.s.s.
- **Maximum entropy principle**
  - Perform extrapolation s.t. the time series characterized by the extrapolated ACF has maximum entropy



- 



## **3.2.2 Maximum Entropy Spectral Estimation (MESE)**

- **Extrapolations of ACF**

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- As long as  $\{r[p+1], r[p+2], \dots\}$  guarantee that the correlation matrix is non-negative definite, they all form valid ACFs for w.s.s.

- **Maximum entropy principle**

- Perform extrapolation s.t. the time series characterized by the extrapolated ACF has maximum entropy
- i.e. the time series will be the least constrained thus most random one among all series having the same first  $(p+1)$  ACF values

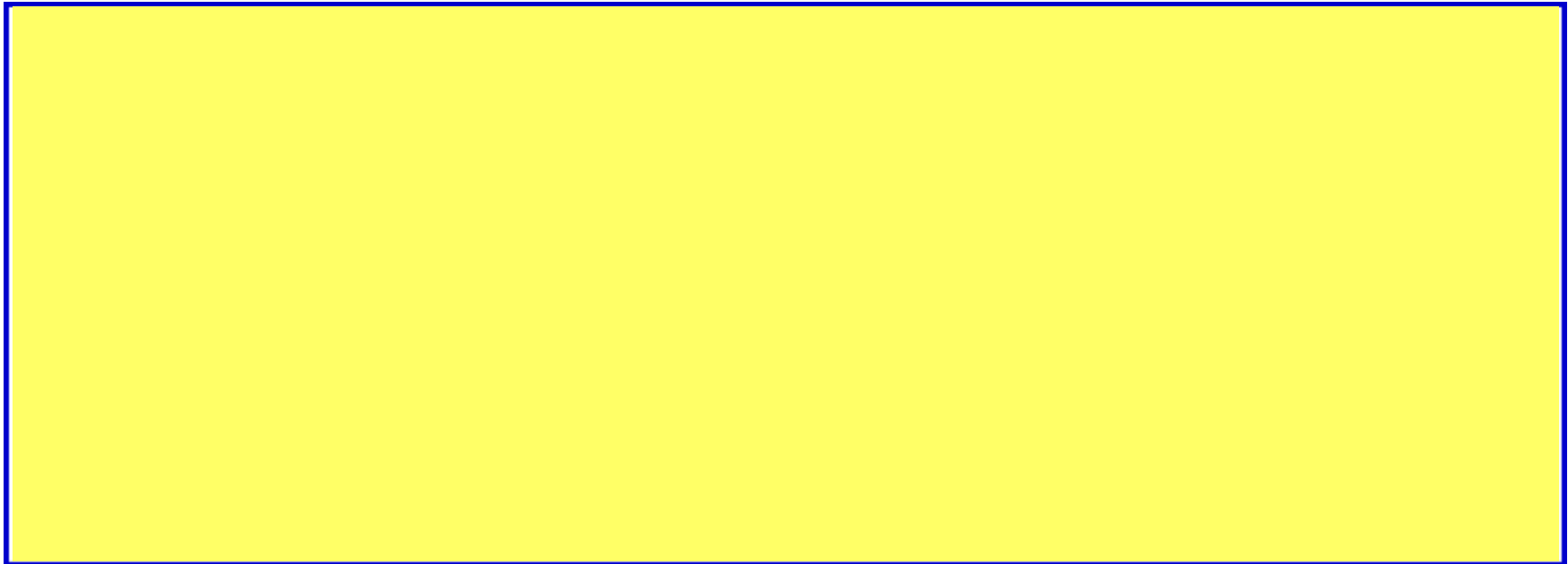
**=> Maximizing entropy leads to estimated p.s.d. be the smoothest one**

- Recall white noise process has flat p.s.d.

## **MESE for Gaussian Process: Formulation**

For a Gaussian random process, the entropy per sample is proportional to

$$\int_{-\frac{1}{2}}^{\frac{1}{2}} \ln P(f) df$$



## **MESE for Gaussian Process: Formulation**

For a Gaussian random process, the entropy per sample is proportional to

$$\int_{-\frac{1}{2}}^{\frac{1}{2}} \ln P(f) df$$

Thus the max entropy spectral estimation is

$$\max \int_{-\frac{1}{2}}^{\frac{1}{2}} \ln P(f) df$$

subject to

$$\int_{-\frac{1}{2}}^{\frac{1}{2}} P(f) e^{j2\pi fk} df = r(k), \quad \text{for } k = 0, 1, \dots, p$$



## **MESE for Gaussian Process: Solution**

Using the Lagrangian multiplier technique, the solution can be found as

$$\hat{P}_{\text{ME}}(f) = \frac{\sigma^2}{\left| 1 + \sum_{k=1}^p a_k e^{-j2\pi f k} \right|^2}$$

## MESE for Gaussian Process: Solution

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$$\hat{P}_{\text{ME}}(f) = \frac{\sigma^2}{\left| 1 + \sum_{k=1}^p a_k e^{-j2\pi f k} \right|^2}$$

where  $\{a_k\}$  are found by solving the Yule-Walker equations given the ACF values  $r(0), \dots, r(p)$

- For Gaussian processes, the MESE is equivalent to AR spectral estimator and the  $P_{\text{ME}}(f)$  is an all-pole spectrum
  - Different assumptions on the process: Gaussian vs AR processes

### 3.2.3 MA Spectral Estimation

An MA(q) model

$$x[n] = \sum_{k=0}^q b_k v[n-k] \quad \Rightarrow \quad B(z) = \sum_{k=0}^q b_k z^{-k}$$

can be used to define an MA spectral estimator

$$\hat{P}_{\text{MA}}(f) = \sigma^2 \left| 1 + \sum_{k=1}^q b_k e^{-j2\pi f k} \right|^2$$

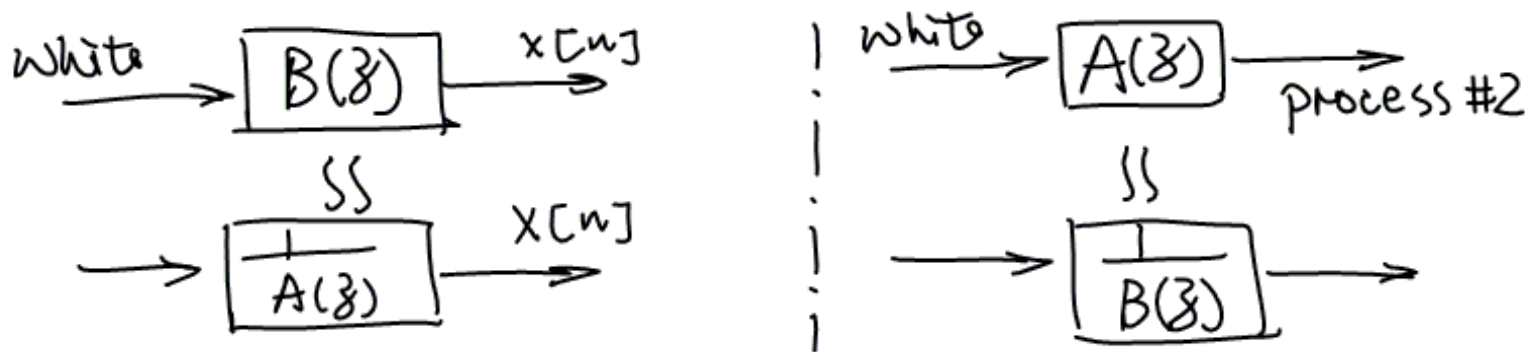
Recall important results on MA process:

- (1) The problem of solving for  $b_k$  given  $\{r(k)\}$  is to solve a set of nonlinear equations;
- (2) An MA process can be approximated by an AR process of sufficiently high order.

## Basic Idea to Avoid Solving Nonlinear Equations

Consider two processes:

- **Process#1:** we observed  $N$  samples, and need to perform spectral estimate
  - We first model it as a high-order AR process, generated by  $1/A(z)$  filter



- **Process#2:** an MA-process generated by  $A(z)$  filter
  - Since we know  $A(z)$ , we can know process#2's autocorrelation function;
  - We model process#2 as an AR( $q$ ) process  $\Rightarrow$  the filter would be  $1/B(z)$

## Use AR Model to Help Finding MA Parameters

- For simplicity, we consider the real coefficients for the MA model.

Note  $P_{MA}(z) = \sigma^2 B(z)B(z^{-1})$

To approximate it with an AR(L) model, i.e.,

$$\text{where } A(z) = 1 + \sum_{k=1}^L a_k z^{-k}$$

$L \gg q$



## Use AR Model to Help Finding MA Parameters

- For simplicity, we consider the real coefficients for the MA model.

Note  $P_{MA}(z) = \sigma^2 B(z)B(z^{-1})$

To approximate it with an AR(L) model, i.e.,

$$P_{MA}(z) \approx \frac{\sigma^2}{A(z)A(z^{-1})} \quad \text{where } A(z) = 1 + \sum_{k=1}^L a_k z^{-k}$$

$L \gg q$

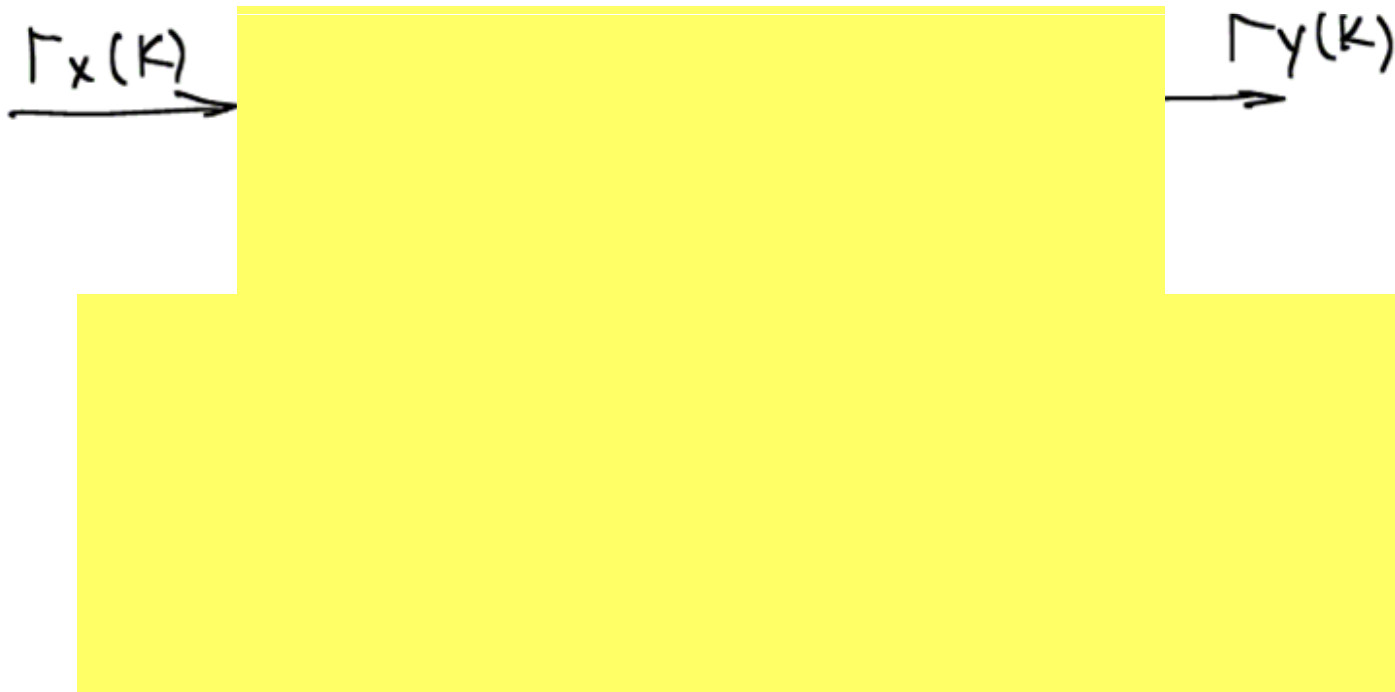
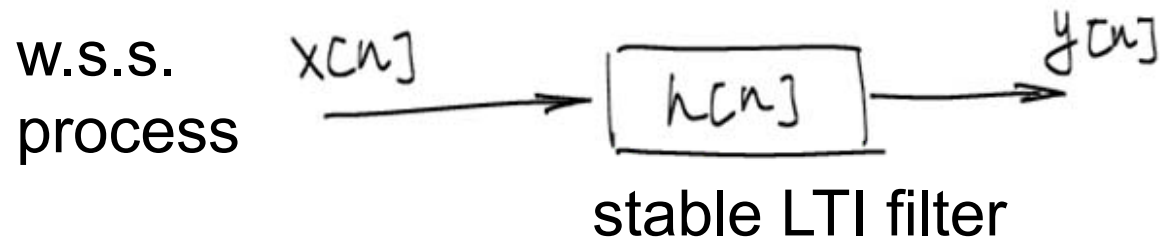
$$\Rightarrow \underbrace{A(z)A(z^{-1})}_{\text{order L}} \approx \frac{1}{\underbrace{B(z)B(z^{-1})}_{\text{order q}}}$$

- ❖ The RHS represents power spectrum of an AR(q) process
- ❖ The unverse ZT of LHS is the ACF of the AR(q) process

# **Use AR Model to Find MA Parameters: Solutions**

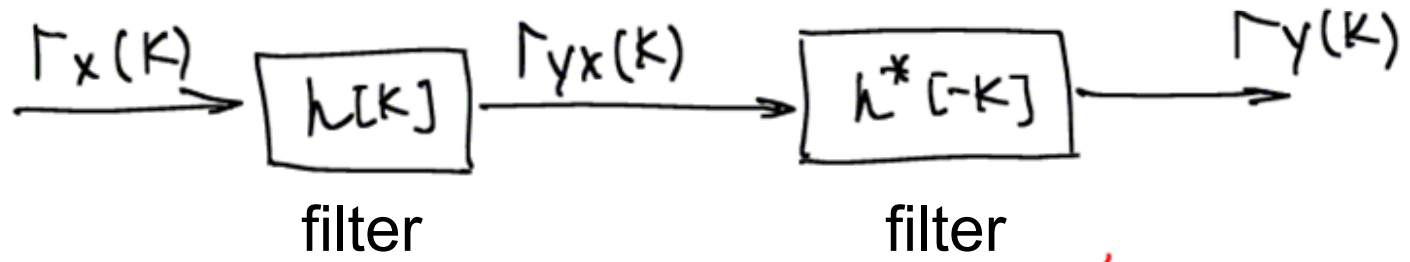
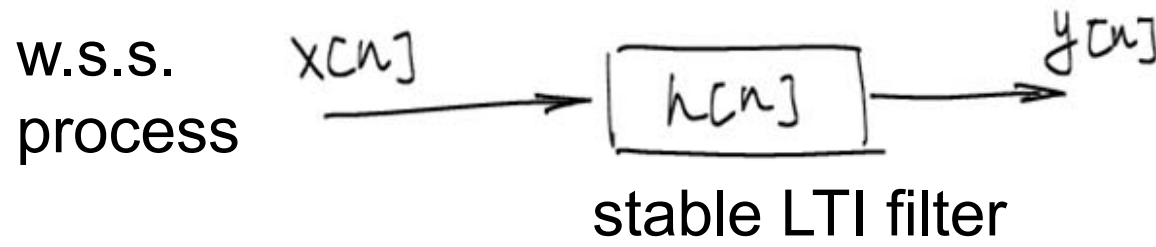
- For simplicity, we consider the real coefficients for the MA model.

## Recall: ACF of Output Process After LTI Filtering





# Recall: ACF of Output Process After LTI Filtering



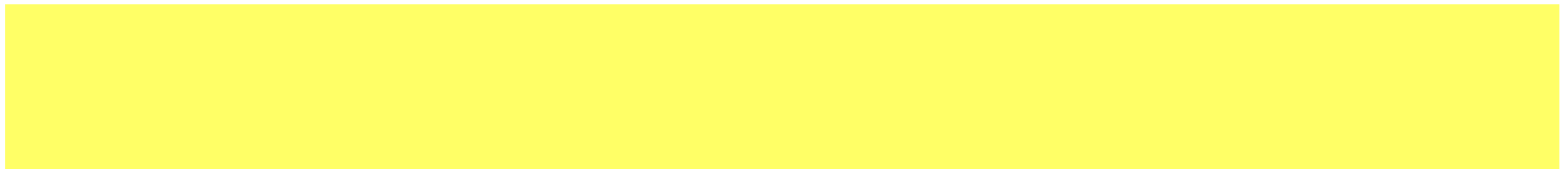
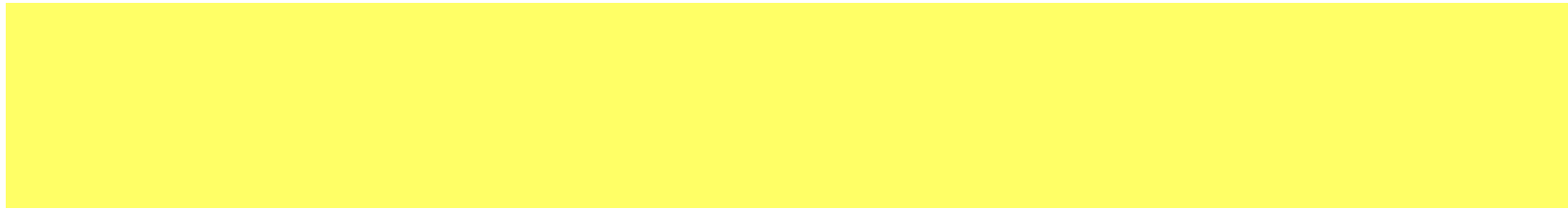
$$\Gamma_h[k] = h[k] * h^*[-k] = \sum_{l=-\infty}^{+\infty} h[l] h^*[k+l]$$

$\downarrow$  ZT  
 $H(z) H^*(1/z^*)$

deterministic autocorrelation of filter's impulse response

## **Use AR to Help Finding MA Parameters (cont'd)**

A random process with power spectrum  $A(z)A(z^{-1})$  can be viewed as filtering a white process by a filter  $A(z)$ , and has autocorrelation



## Use AR to Help Finding MA Parameters (cont'd)

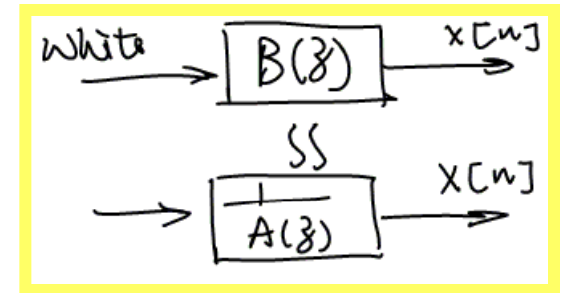
A random process with power spectrum  $A(z)A(z^{-1})$  can be viewed as filtering a white process by a filter  $A(z)$ , and has autocorrelation

proportional to  $\sum_{n=0}^{L-k} a_n a_{n+k}$  for lag  $k$

→ Knowing such autocorrelation function, we can use Levinson-Durbin recursion to find the optimal linear prediction parameters for the process (or equivalently its AR approximation parameters)

Thus we get  $\{b_k\}$  as  $A(z)A(z^{-1}) \approx \frac{1}{B(z)B(z^{-1})}$

# Durbin's Method



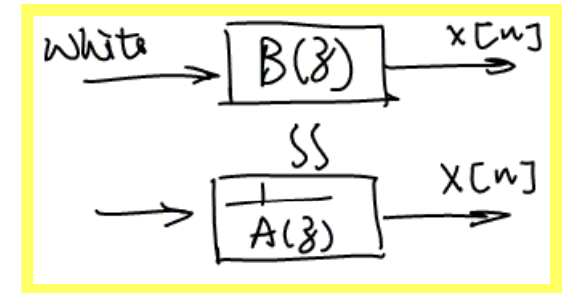
1. Use Levinson-Durbin recursion and solve for

$$\begin{bmatrix} \hat{\Gamma}(0) & \hat{\Gamma}(1) & \dots & \hat{\Gamma}(L-1) \\ \hat{\Gamma}(1) & \hat{\Gamma}(0) & & \hat{\Gamma}(L-2) \\ \vdots & & \ddots & \vdots \\ \hat{\Gamma}(L-1) & \dots & \dots & \hat{\Gamma}(0) \end{bmatrix} \begin{bmatrix} \hat{a}_1 \\ \hat{a}_2 \\ \vdots \\ \hat{a}_L \end{bmatrix} = - \begin{bmatrix} \hat{\Gamma}(1) \\ \vdots \\ \hat{\Gamma}(L) \end{bmatrix}$$

where

$$\hat{\Gamma}(k) = \frac{1}{N} \sum_{n=0}^{N-1-k} x[n] x[n+k]$$

# Durbin's Method



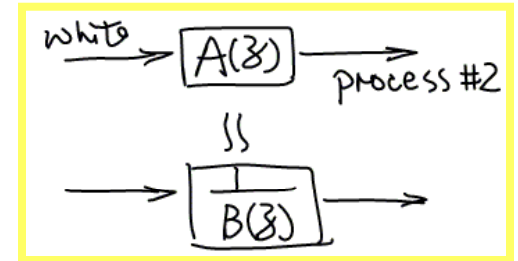
1. Use Levinson-Durbin recursion and solve for

$$\begin{bmatrix} \hat{r}(0) & \hat{r}(1) & \dots & \hat{r}(L-1) \\ \hat{r}(1) & \hat{r}(0) & & \hat{r}(L-2) \\ \vdots & & \ddots & \vdots \\ \hat{r}(L-1) & \dots & \dots & \hat{r}(0) \end{bmatrix} \begin{bmatrix} \hat{a}_1 \\ \hat{a}_2 \\ \vdots \\ \hat{a}_L \end{bmatrix} = - \begin{bmatrix} \hat{r}(1) \\ \vdots \\ \hat{r}(L) \end{bmatrix}$$

where 
$$\hat{r}(k) = \frac{1}{N} \sum_{n=0}^{N-1-k} x[n] x[n+k]$$

- We first approximate the observed data sequence  $\{x[0], \dots, x[N]\}$  with an **AR model** of high order (often pick  $L > 4q$ )
- We use **biased ACF estimate** here to ensure nonnegative definiteness and smaller variance than unbiased estimate (dividing by  $N-k$ )

## Durbin Method (cont'd)

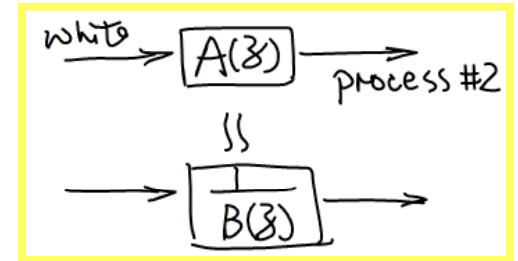


2. Fit the data sequence  $\{1, \hat{a}_1, \hat{a}_2, \dots, \hat{a}_L\}$  to an AR(q) model:

$$\begin{bmatrix} \hat{\Gamma}_a(0) & \hat{\Gamma}_a(1) & \dots & \hat{\Gamma}_a(q-1) \\ \hat{\Gamma}_a(1) & \hat{\Gamma}_a(0) & & \hat{\Gamma}_a(q-2) \\ \vdots & & \ddots & \vdots \\ \hat{\Gamma}_a(q-1) & \dots & \dots & \hat{\Gamma}_a(0) \end{bmatrix} \begin{bmatrix} \hat{b}_1 \\ \vdots \\ \hat{b}_q \end{bmatrix} = - \begin{bmatrix} \hat{\Gamma}_a(1) \\ \vdots \\ \hat{\Gamma}_a(q) \end{bmatrix}$$

- Note we add  $1/(L+1)$  factor to allow the interpretation of  $r_a(k)$  as an autocorrelation function estimator

## Durbin Method (cont'd)



2. Fit the data sequence  $\{1, \hat{a}_1, \hat{a}_2, \dots, \hat{a}_L\}$  to an AR(q) model:

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where  $\hat{\Gamma}_a(k) = \frac{1}{L+1} \sum_{n=0}^{L-k} \hat{a}_n \hat{a}_{n+k}$

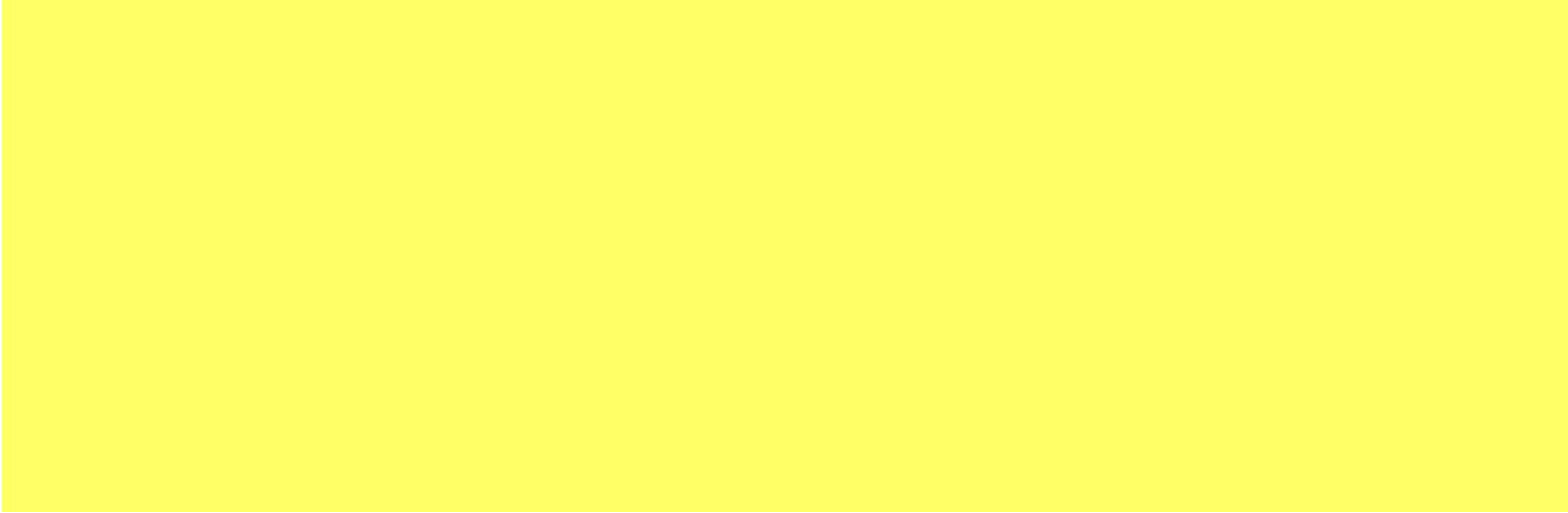
- The result  $\{b_i\}$  is the estimated MA parameters for original  $\{x[n]\}$
- Note we add  $1/(L+1)$  factor to allow the interpretation of  $r_a(k)$  as an autocorrelation function estimator

## **3.2.4 ARMA Spectral Estimation**

Recall the ARMA(p,q) model



We define an ARMA(p,q) spectral estimator





## 3.2.4 ARMA Spectral Estimation

Recall the ARMA(p,q) model

$$x[n] = -\sum_{k=1}^p a_k x[n-k] + \sum_{k=0}^q b_k v[n-k]$$

We define an ARMA(p,q) spectral estimator

$$\hat{P}_{\text{ARMA}}(f) = \hat{\sigma}^2 \frac{\left| 1 + \sum_{k=1}^q \hat{b}_k e^{-j2\pi f k} \right|^2}{\left| 1 + \sum_{k=1}^p \hat{a}_k e^{-j2\pi f k} \right|^2}$$

## Modified Yule-Walker Equations

Recall the Yule-Walker Eq. for ARMA(p,q) process

$$\left\{ \begin{array}{l} \Gamma_x[k] = - \sum_{l=1}^p a[l] \Gamma_x[k-l] + \sigma^2 \sum_{l=0}^{q-k} h^*[l] b[l+k] \\ \Gamma_x[k] = - \sum_{l=1}^p a[l] \Gamma_x[k-l] \end{array} \right. , \quad k=0,1,\dots,q$$
$$\Gamma_x[k] = - \sum_{l=1}^p a[l] \Gamma_x[k-l] \quad , \quad k \geq q+1 .$$

## Modified Yule-Walker Equations

Recall the Yule-Walker Eq. for ARMA(p,q) process

$$\begin{cases} \Gamma_x[k] = - \sum_{l=1}^p a[l] \Gamma_x[k-l] + \sigma^2 \sum_{l=0}^{q-k} h^*[l] b[l+k] & k=0,1,\dots,q \\ \Gamma_x[k] = - \sum_{l=1}^p a[l] \Gamma_x[k-l] & , k \geq q+1. \end{cases}$$

We can use the equations for  $k \geq q+1$  to solve for  $\{a_i\}$

$$\begin{bmatrix} \Gamma(q) & \Gamma(q-1) & \dots & \Gamma(q-p+1) \\ \Gamma(q+1) & \Gamma(q) & & \vdots \\ \vdots & & \ddots & \vdots \\ \Gamma(q+p-1) & \dots & \dots & \Gamma(q) \end{bmatrix} \begin{bmatrix} a_1 \\ a_2 \\ \vdots \\ a_p \end{bmatrix} = - \begin{bmatrix} \Gamma(q+1) \\ \vdots \\ \Gamma(q+p) \end{bmatrix}$$

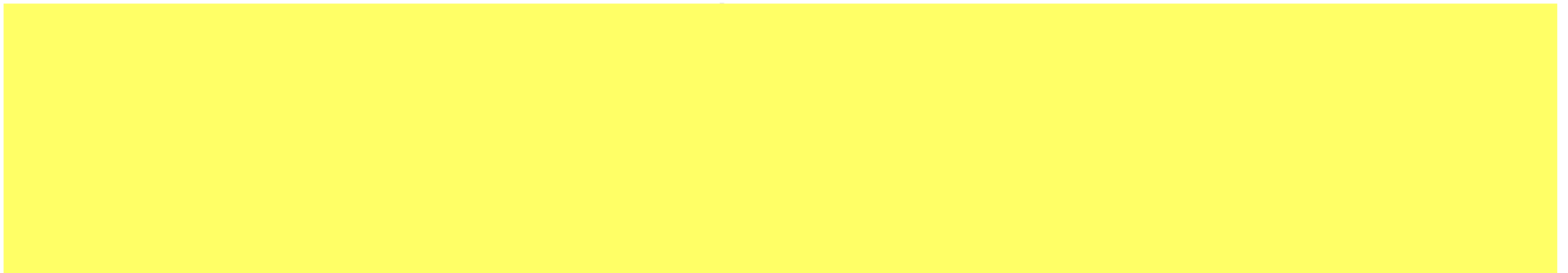
$$\Rightarrow S \hat{\underline{a}} = \underline{t} \quad \text{“Modified Yule-Walker Equations”}$$

## Estimating ARMA Parameters

1. By solving the modified Yule-Walker eq., we get

$$\hat{A}(z) = 1 + \sum_{k=1}^p \hat{a}_k z^{-k}$$

2. To estimate  $\{b_k\}$ ,



## Estimating ARMA Parameters

1. By solving the modified Yule-Walker eq., we get

$$\hat{A}(z) = 1 + \sum_{k=1}^p \hat{a}_k z^{-k}$$

2. To estimate  $\{b_k\}$ ,

We first filter  $\{x[n]\}$  with  $\hat{A}(z)$ , and model the output with an MA(q) model using Durbin's method.

## Extension: LSMYWE Estimator

- Performance by solving  $p$  modified Yule-Walker equations followed by Durbin's method
    - May yield highly variable spectral estimates (esp. when the matrix involving ACF is nearly singular due to poor ACF estimates)
  - Improvement: use more than  $p$  equations to solve  $\{a_1 .. a_p\}$  in a least square sense
    - Use Yule-Walker equations for  $k = (q+1), \dots, M$ :  $\min || \underline{t} - \underline{S} \underline{a} ||^2$
    - Least square solution:  $\underline{a} = (\underline{S}^H \underline{S})^{-1} \underline{S}^H \underline{t}$
    - Then obtain  $\{b_i\}$  by Durbin's method
- ➔ “Least-square modified Yule-Walker equation” (LSMYWE)

Ref: review in Hayes' book Sec.2.3.6 on least square solution

## Comparison of Different Methods: Revisit

- Test case: a process consists of narrowband components (sinusoids) and a broadband component (AR)

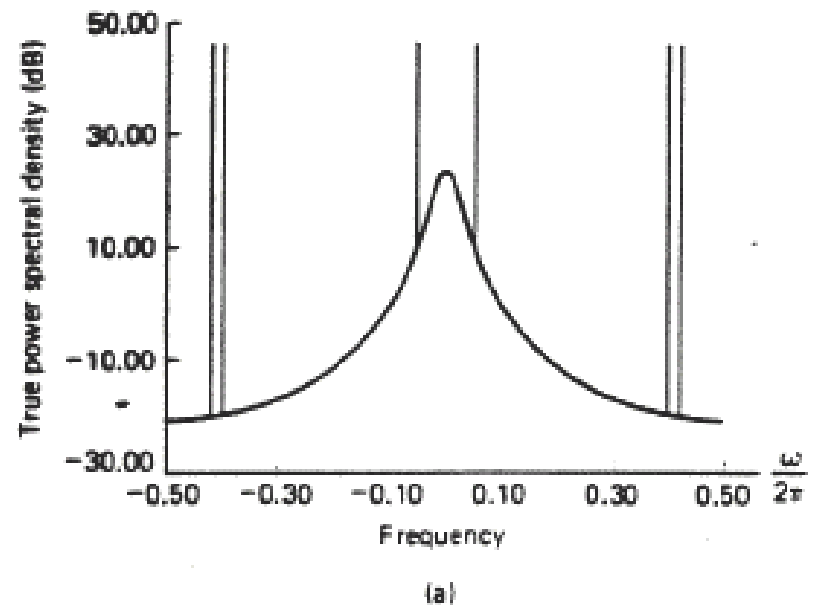
- $x[n] = 2 \cos(\omega_1 n) + 2 \cos(\omega_2 n) + 2 \cos(\omega_3 n) + z[n]$

- where  $z[n] = -a_1 z[n-1] + v[n]$ ,  $a_1 = -0.85$ ,  $\sigma^2 = 0.1$

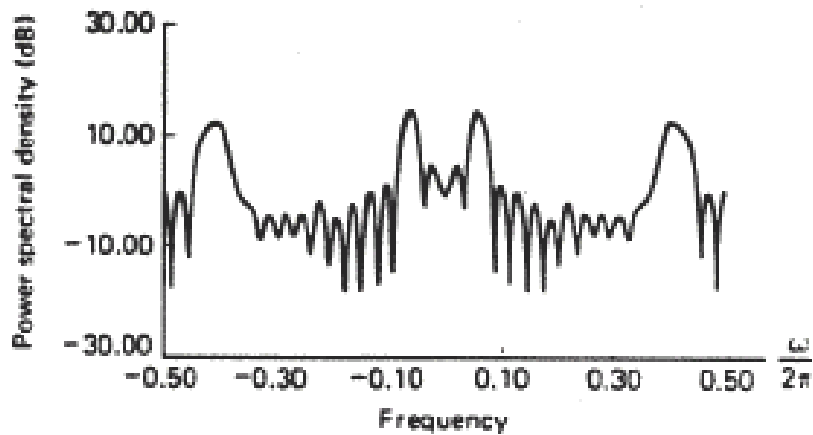
- $\omega_1/2\pi = 0.05$ ,  $\omega_2/2\pi = 0.40$ ,  $\omega_3/2\pi = 0.42$

- $N=32$  data points are available  
→ periodogram resolution  $f = 1/32$

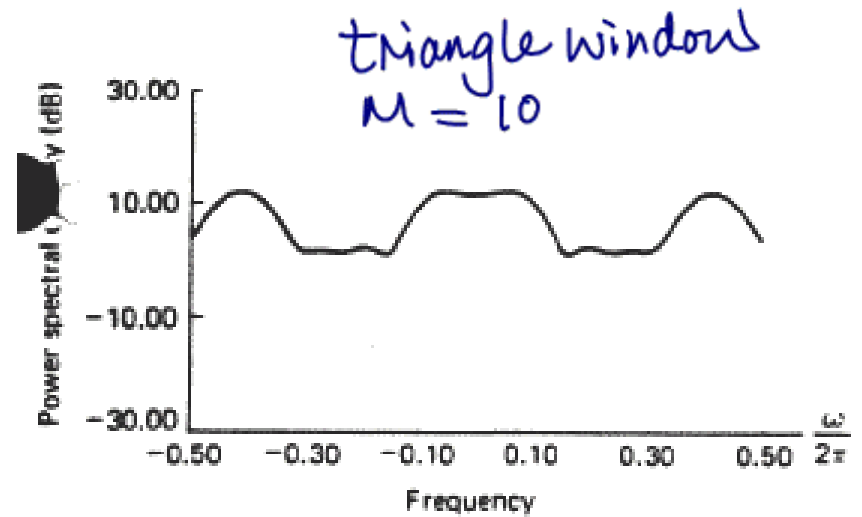
- Examine typical characteristics of various non-parametric and parametric spectral estimators



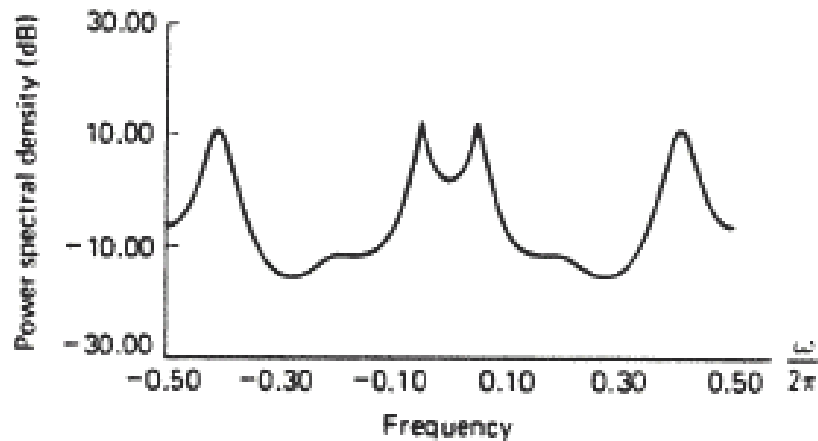
(Fig.2.17 from Lim/Oppenheim book)



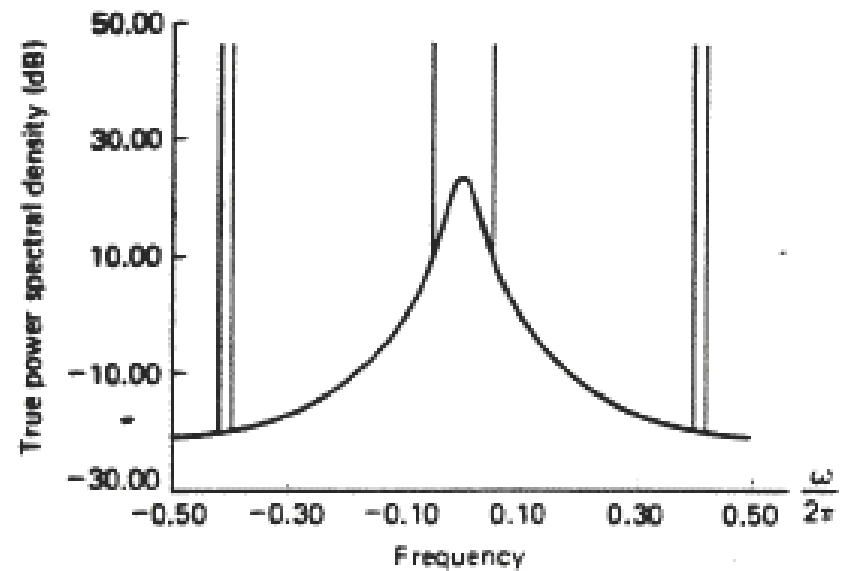
(b) Periodogram



(c) Blackman-Tukey



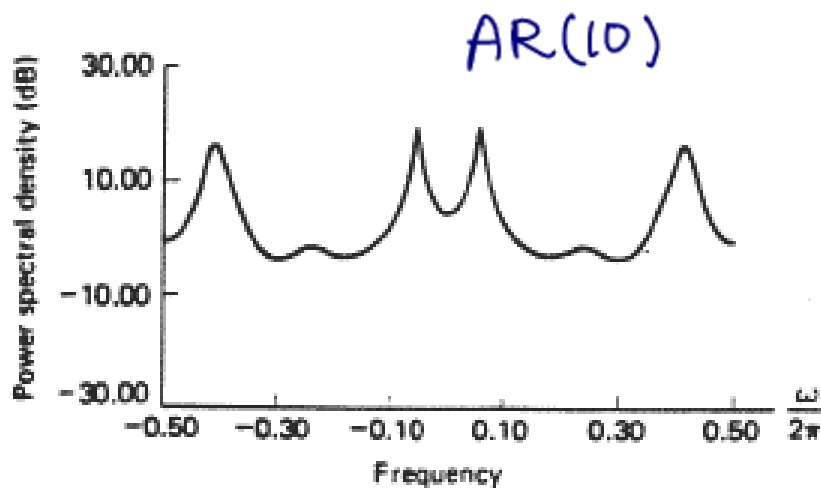
(d) Minimum variance spectral estimator



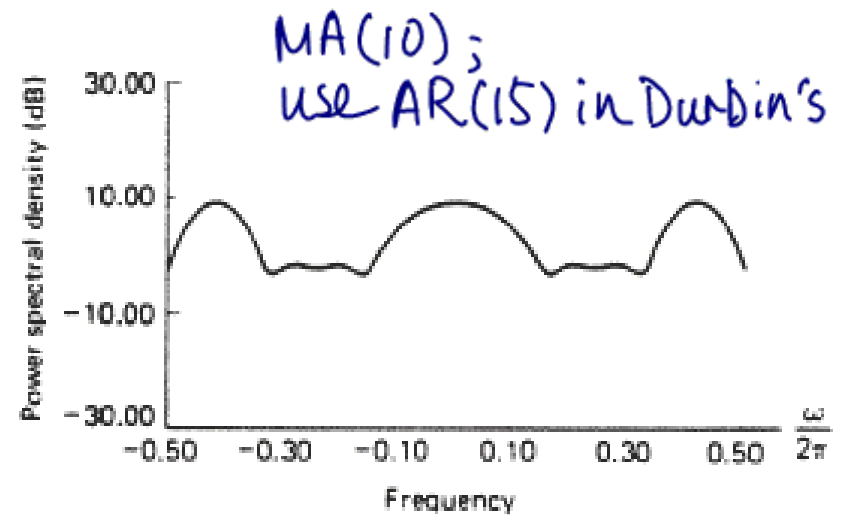
true p.s.d.

(a)

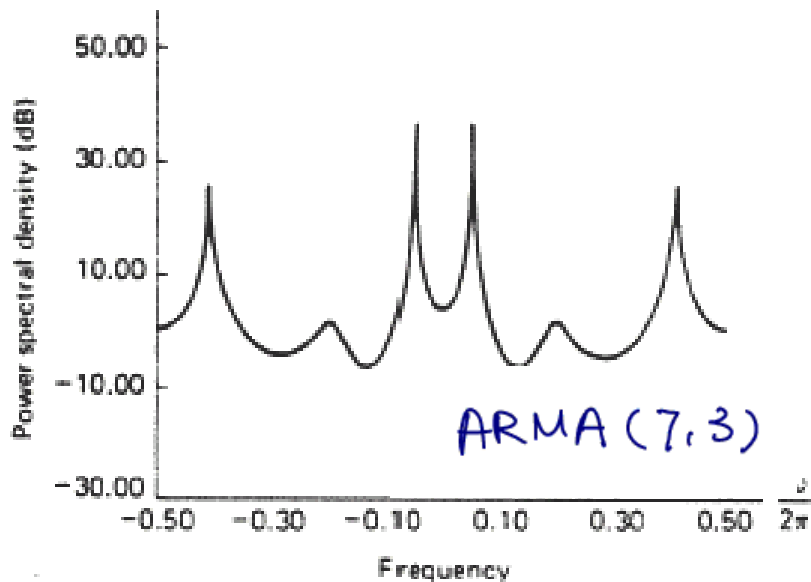




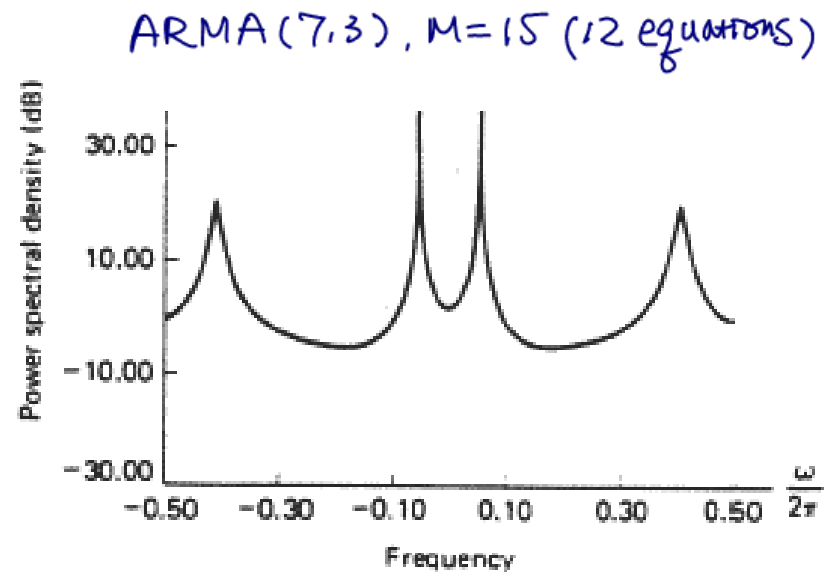
(e) Autocorrelation



(i) Durbin



(j) Modified Yule-Walker equations



(k) Least-squares modified Yule-Walker equations

## **3.2.5 Model Order Selection**

- The best way to determine the model order is to base it on the physics of the data generation process
- Example: speech processing
  - Studies show the vocal tract can be modeled as an all-pole filter having 4 resonances in a 4kHz band, thus at least 4 pairs of complex conjugate poles are necessary
    - ➔ Typically 10-12 poles are used in an AR modeling for speech
- When no such knowledge is available, we can use some **statistical test** to estimate the order

Ref. for in-depth exploration: “Model-order selection,” by P. Stoica and Y. Selen, IEEE Signal Processing Magazine, July 2004.

# Considerations for Order Selection

- Modeling error
  - Modeling error measures the (statistical) difference between the true data value and the approximation by the model
    - ◆ *e.g. estimating linear prediction MSE in AR modeling*
  - Usually for a given type of model (e.g. AR, ARMA), the modeling error decreases as we increase the model order
- Balance between the modeling error and the amount of model parameters to be estimated
  - The number of parameters that need to be estimated and represented increases as we use higher model order → Cost of overmodeling
  - Can balance modeling error and the cost of going to higher model by imposing a penalty term that increases with the model order

# **A Few Commonly Used Criteria**

- **Akaike Information Criterion (AIC)**

- A general estimate of the Kullback-Leibler divergence between assumed and true p.d.f., with an order penalty term increasing linearly
- Choose the model order that minimize AIC

- **Minimum Description Length (MDL) Criterion**

- Impose a bigger penalty term to overcome AIC's overestimation
- Estimated order converges to the true order as  $N$  goes to infinity

# A Few Commonly Used Criteria

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- Choose the model order that minimize AIC

$$\text{AIC}(i) = N \ln \varepsilon_p + 2i$$

size of  
available data

model error

model order:  
 $i=p$  for AR( $p$ )  
 $i=p+q$  for ARMA( $p,q$ )

- Minimum Description Length (MDL) Criterion

- Impose a bigger penalty term to overcome AIC's overestimation
- Estimated order converges to the true order as  $N$  goes to infinity

$$\text{MDL}(i) = N \ln \varepsilon_p + (\log N)i$$