ENEE 630 Homework 8^1

Material Covered: Lattice Predictor, Reflection Coefficient, Non-parametric Spectrum Estimation and ACF estimator

Problem 1 In a normalized lattice predictor, the forward and backward prediction errors at the various stage of the predictor are all normalized to have unit variance. Such an operation makes it possible to utilize the full dynamic range of multipliers used in the predictor, the normalized forward and backward prediction errors are defined as follows, respectively:

$$\overline{f}_m(n) = \frac{f_m(n)}{P_m^{1/2}}$$
$$\overline{b}_m(n) = \frac{b_m(n)}{P_m^{1/2}}$$

where P_m is the average power (or variance) of the forward prediction error $f_m(n)$ or that of the backward prediction error $b_m(n)$. Show that the structure of stage m of the normalized lattice predictor is as shown in Figure 1 for real-valued data.

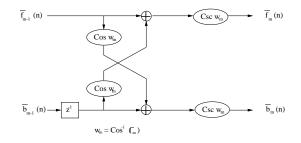


Figure 1: P-1

Problem 2 Consider the problem of optimizing stage m of the lattice predictor. The cost function to be used in the optimization is described by

$$J_m(\Gamma_m) = \alpha E[|f_m(n)|^2] + (1 - \alpha)E[|b_m(n)|^2]$$

where α is a constant that lies between zero and one; $f_m(n)$ and $b_m(n)$ denote the forward and backward prediction errors at the output of stage m, respectively.

 $^{^{1}}$ ver. 201211

(a) Show that the optimum value of the reflection coefficient Γ_m for which J_m is at minimum equals

$$\Gamma_{m,o}(\alpha) = -\frac{E[b_{m-1}(n-1)f_{m-1}^*(n)]}{(1-\alpha)E[|f_{m-1}(n)|^2] + \alpha E[|b_{m-1}(n-1)|^2]}$$

- (b) Evaluate $\Gamma_{m,o}(\alpha)$ for each of the following three special conditions:
 - (1) α = 1
 (2) α = 0
 - (3) $\alpha = \frac{1}{2}$

Notes: When the parameter $\alpha = 1$, the cost function reduces to

$$J_m(\Gamma_m) = E[|f_m(n)|^2]$$

We refer to this criterion as the *forward method*. When the parameter $\alpha=0$, the cost function reduces to

$$J_m(\Gamma_m) = E[|b_m(n)|^2]$$

We refer to this criterion as the *backward method*.

When the parameter $\alpha = \frac{1}{2}$, the formula for $\Gamma_{m,o}(\alpha)$ reduces to what is known as the *Burg formula*.

Problem 3 In this problem we show that the periodogram is an inconsistent estimator by examining the estimator at f = 0, or

$$\hat{P}_{PER}(0) = \frac{1}{N} \left(\sum_{n=0}^{N-1} x[n] \right)^2$$

If x[n] is a real white Gaussian noise process with PSD

$$P_{xx}(f) = \sigma_x^2$$

find the mean and variance of $\hat{P}_{PER}(0)$. Does the variance converge to zero as $N \to \infty$? Hint: Note that

$$\hat{P}_{PER}(0) = \sigma_x^2 \left(\sum_{n=0}^{N-1} \frac{x[n]}{\sigma_x \sqrt{N}}\right)^2$$

where the quantity inside the parentheses is $\sim N(0, 1)$.

Problem 4 Consider the estimator

$$\hat{P}_{AVPER}(0) = \frac{1}{N} \sum_{m=1}^{N-1} \hat{P}_{PER}^{(m)}(0)$$

where

$$\hat{P}_{PER}^{(m)}(0) = x^2[m]$$

for the process from Problem 3. This estimator may be viewed as an *averaged periodogram*. In essence the data record is sectioned into blocks(in this case, of length 1) and the periodograms for each block are averaged. Find the mean and variance of $\hat{P}_{AVPER}(0)$. Compare this result to that obtained in Problem 3.

Problem 5 For the real data set $\{x[0], x[1], x[2]\}$ the unbiased ACF estimator yields

$$\hat{r}'_{xx}[0] = \frac{1}{3}(x^2[0] + x^2[1] + x^2[2])$$
$$\hat{r}'_{xx}[1] = \frac{1}{2}(x[0]x[1] + x[1]x[2])$$

Show that the autocorrelation sequence is not positive semidefinite by giving a numerical example where the matrix

$$\begin{bmatrix} \hat{r}'_{xx}[0] & \hat{r}'_{xx}[1] \\ \hat{r}'_{xx}[1] & \hat{r}'_{xx}[0] \end{bmatrix}$$

is not positive semidefinite.